

Name: \_\_\_\_\_ ND ID: \_\_\_\_\_ Date: \_\_\_\_\_

Signature: \_\_\_\_\_.

(By signing here, I certify that I have taken this test while refraining from cheating.)

## Mid-Term 1

This exam contains 8 pages (including this cover page) and 5 problems. Check to see if any pages are missing. Enter all requested information on the top of this page.

You may *not* use your book or any calculator on this exam. You *can* use your homeworks. You are required to show your work on each problem on the exam. The following rules apply:

- You have 60 minutes to complete the exam, starting at the beginning of class.
- **If you use a theorem or proposition from class or the notes or the book you must indicate this** and explain why the theorem may be applied. It is okay to just say, “by some theorem/proposition from class.”
- **Organize your work**, in a reasonably neat and coherent way, in the space provided. Work scattered all over the page without a clear ordering will receive very little credit.
- **Mysterious or unsupported answers will not receive full credit.** A correct answer, unsupported by calculations, explanation, or algebraic work will receive no credit; an incorrect answer supported by substantially correct calculations and explanations might still receive partial credit.
- If you need more space, use the back of the pages; clearly indicate when you have done this. Scratch paper appears at the end of the document.

Problem	Points	Score
1	30	
2	30	
3	30	
4	30	
5	30	
Total:	150	

Do not write in the table to the right. Good luck!<sup>a</sup>

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## Reference sheet

Below are some definitions that may be relevant.

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Let  $\Omega$  be a universal set and let  $\mathcal{F}$  be a  $\sigma$ -algebra on  $\Omega$ . A **measure**  $\mu$  is a function  $\mu: \mathcal{F} \rightarrow [0, \infty]$  such that (i)  $\mu(\emptyset) = 0$ , and (ii) if  $A_1, A_2, \dots \in \mathcal{F}$  are disjoint, then  $\mu(\cup_{i=1}^{\infty} A_i) = \sum_{i=1}^{\infty} \mu(A_i)$ . A **probability measure** on  $\Omega$  is a measure  $\mu$  such that  $\mu(\Omega) = 1$ .

Let  $X: \Omega \rightarrow S$  be measurable. The **distribution** of  $X$  (or the **law** of  $X$ ) is the probability measure  $\mu_X$  defined for any measurable  $A \subseteq S$  by

$$\mu_X(A) := \mathbf{P}(X \in A) = \mathbf{P}(\{\omega \in \Omega: X(\omega) \in A\}).$$

We say that a sequence of random variables  $X_1, X_2, \dots: \Omega \rightarrow \mathbf{R}$  **converges in probability** to a random variable  $X: \Omega \rightarrow \mathbf{R}$  if: for all  $\varepsilon > 0$

$$\lim_{n \rightarrow \infty} \mathbf{P}(|X_n - X| > \varepsilon) = 0.$$

We say that a sequence of real-valued random variables  $X_1, X_2, \dots$  **converges in distribution** to a real-valued random variable  $X$  if, for any  $t \in \mathbf{R}$  such that  $\mathbf{P}(X \leq t)$  is continuous at  $t$ ,

$$\lim_{n \rightarrow \infty} \mathbf{P}(X_n \leq t) = \mathbf{P}(X \leq t).$$

We say that a sequence of random variables  $X_1, X_2, \dots: \Omega \rightarrow \mathbf{R}$  **converges in  $L_2$**  to a random variable  $X: \Omega \rightarrow \mathbf{R}$  if

$$\lim_{n \rightarrow \infty} \mathbf{E}|X_n - X|^2 = 0.$$

We say that a sequence of random variables  $X_1, X_2, \dots: \Omega \rightarrow \mathbf{R}$  **converges almost surely** to a random variable  $X: \Omega \rightarrow \mathbf{R}$  if

$$\mathbf{P}(\lim_{n \rightarrow \infty} X_n = X) = 1.$$

Let  $X: \Omega \rightarrow [-\infty, \infty]$  be a random variable with  $\mathbf{E}|X| < \infty$  and  $\mathbf{E}X^2 < \infty$ . We define the **variance** of  $X$ , denoted  $\text{var}(X)$ , to be  $\text{var}(X) := \mathbf{E}(X - \mathbf{E}X)^2 = \mathbf{E}X^2 - (\mathbf{E}X)^2$ .

We say a collection  $(X_i: \Omega \rightarrow S_i)_{i \in I}$  of random variables is **independent** if the distribution of  $(X_i)_{i \in I}$  is the product of the distributions of the  $X_i$ . That is, for any finite  $J \subseteq I$  and for any measurable sets  $A_i \subseteq S_i$ ,  $i \in J$ , we have

$$\mathbf{P}\left(\bigcap_{i \in J} \{X_i \in A_i\}\right) = \prod_{i \in J} \mathbf{P}(X_i \in A_i).$$

1. (a) (15 points) Let  $X: \Omega \rightarrow \mathbf{R}$  be a random variable such that

$$\mathbf{P}(X > 0) > 0.$$

Show that there exists a positive integer  $n$  such that  $\mathbf{P}(X > 1/n) > 0$ .

- (b) (15 points) Let  $X_1, X_2, \dots \Omega \rightarrow \mathbf{R}$  be random variables such that  $\mathbf{E}X_i = 0$  and  $\mathbf{E}X_i^2 = 1$  for all  $i \geq 1$ . Show that

$$\mathbf{P}(X_n > n \text{ for infinitely many } n \geq 1) = 0.$$

2. (30 points) Let  $X_1, X_2, \dots$  be independent identically distributed random variables with  $\mathbf{P}(X_1 = 1) = \mathbf{P}(X_1 = -1) = 1/2$ . Let  $a_1, a_2, \dots \in \mathbf{R}$ . Show that, for any  $n \geq 1$ ,

$$\mathbf{P}\left(\sum_{i=1}^n a_i X_i \geq t\right) \leq e^{-\frac{t^2}{2\sum_{i=1}^n a_i^2}}, \quad \forall t \geq 0.$$

(You can use the following inequality without proof:  $\cosh(x) \leq e^{x^2/2}, \forall x \in \mathbf{R}$ .)

3. (30 points) Let  $n$  be a positive integer. Suppose  $X_1, X_2, \dots$  are independent random variables that are uniformly distributed in the set  $\{1, 2, \dots, n\}$ . We can think of  $\{1, 2, \dots, n\}$  as a set of baseball cards, and for any  $i \geq 1$ ,  $X_i$  is a uniformly random baseball card that you have found. Your goal is to collect all of the  $n$  baseball cards.

For any  $0 \leq j \leq n$ , let  $T_j$  be the first time that you have collected exactly  $j$  baseball cards. That is,  $T_j$  is the smallest integer  $k$  such that  $\{X_1, \dots, X_k\}$  consists of exactly  $k$  distinct elements. For example,  $T_0 = 0$ ,  $T_1 = 1$ ,  $T_2$  is 2 when  $X_2 \neq X_1$ ,  $T_2$  is 3 when  $X_2 = X_1$  and  $X_3 \neq X_1$ , and so on. (You may assume that  $\mathbf{P}(T_j < \infty) = 1$  for all  $0 \leq j \leq n$ .)

For any  $1 \leq j \leq n$ , let  $Y_j := T_j - T_{j-1}$  be the time it takes to go from a collection of  $j - 1$  distinct baseball cards to a collection of  $j$  distinct baseball cards.

Show that  $Y_2$  and  $Y_3$  are independent, geometric random variables with parameters  $\frac{n-1}{n}$  and  $\frac{n-2}{n}$ , respectively.

(Recall that a geometric random variable  $Y$  with parameter  $0 < p < 1$  is a positive-integer valued random variable such that  $\mathbf{P}(Y = k) = (1 - p)^{k-1}p$  for any  $k \geq 1$ .)

4. (30 points) We continue the definitions and notation from the previous problem. In this problem, you may assume that  $Y_1, \dots, Y_n$  are independent random variables and for any  $1 \leq j \leq n$ ,  $Y_j$  is a geometric random variable with parameter  $\frac{n-j+1}{n}$ .

You may use the following fact: a geometric random variable  $Y$  with parameter  $0 < p < 1$  has mean  $\frac{1}{p}$  and variance  $\frac{1-p}{p}$ .

Note that  $T_n = Y_1 + \dots + Y_n$ .

- Show that  $\mathbf{E}T_n = n \log n + O(n)$  and  $\text{var}(T_n) = O(n^2)$ .
- Conclude that

$$\frac{T_n}{n \log n}$$

converges in probability to 1 as  $n \rightarrow \infty$ .

(Hint: Can you bound  $\mathbf{P}(|T_n - n \log n + O(n)| > tn)$ ?)

So, if you want to complete a set of 100 distinct baseball cards, you would need to randomly sample about  $100 \log 100 \approx 460$  baseball cards.

(As usual,  $O(a)$  denotes any quantity whose absolute value is bounded by a constant multiple  $ca$  of  $a$ .)

5. (30 points) Find a sequence of random variables  $X_1, X_2, \dots : \Omega \rightarrow \mathbf{R}$  such that

- $X_1, X_2, \dots$  converges in probability to 0.
- $X_1, X_2, \dots$  does **not** converge almost surely to 0.
- $X_1, X_2, \dots$  does **not** converge in  $L_2$  to 0.

Prove that your example of  $X_1, X_2, \dots$  satisfies the above three properties.

(As usual, it might be easiest to use  $\Omega = [0, 1]$  with  $\mathbf{P}$  Lebesgue measure on  $\Omega$ .)

(Scratch paper)