

## 171 Midterm 1 Solutions, Fall 2016<sup>1</sup>

### 1. QUESTION 1

True/False

(a) Let  $\Omega$  be a universe. Let  $A_1, A_2, \dots \subseteq \Omega$ . Then

$$\bigcup_{i=1}^{\infty} A_i = \{x \in \Omega : \forall \text{ positive integers } j, x \in A_j\}.$$

FALSE. If  $A_1 = \Omega$  and  $A_2 = \emptyset$ , then  $\bigcup_{i=1}^{\infty} A_i = \Omega$ , but  $\emptyset = \{x \in \Omega : \forall \text{ positive integers } j, x \in A_j\}$ .

(b) For any positive integers  $i, j$ , let  $a_{ij}$  be a real number. Then

$$\sum_{i=1}^{\infty} \left( \sum_{j=1}^{\infty} a_{ij} \right) = \sum_{j=1}^{\infty} \left( \sum_{i=1}^{\infty} a_{ij} \right).$$

FALSE. For any  $i \geq 1$ , let  $a_{i(i+1)} = 1$ , let  $a_{ii} = -1$ , and let  $a_{ij} = 0$  for any other  $i, j$ . Then  $\sum_{i=1}^{\infty} (a_{ii} + a_{i(i+1)}) = \sum_{i=1}^{\infty} (0) = 0 \neq -1 = a_{11} + 0 = a_{11} + \sum_{j=2}^{\infty} (a_{(j-1)j} + a_{jj}) = \sum_{j=1}^{\infty} (\sum_{i=1}^{\infty} a_{ij})$

(c) The Markov Chain with transition matrix  $P = \begin{pmatrix} 0 & 1 & 0 \\ 0 & 0 & 1 \\ 1 & 0 & 0 \end{pmatrix}$  has exactly two recurrent states.

FALSE; All three states are recurrent. Since  $P^3 = \begin{pmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{pmatrix}$ , for any  $x \in \{1, 2, 3\}$  we have  $\mathbf{P}_x(X_3 = x) = 1$ . So,  $\mathbf{P}_x(T_x \leq 3) = 1$ , and  $\mathbf{P}_x(T_x < \infty) = 1$ .

### 2. QUESTION 2

Suppose  $X$  and  $Y$  are independent standard Gaussian distributed random variables. (So,  $\mathbf{P}(a \leq X \leq b) = \int_a^b e^{-t^2/2} dt / \sqrt{2\pi}$ , for any  $-\infty \leq a \leq b \leq \infty$ .) Find the probability density function of  $X + Y$ . That is, find a function  $f_{X+Y}: \mathbb{R} \rightarrow [0, \infty)$  such that  $\mathbf{P}(a \leq X + Y \leq b) = \int_a^b f_{X+Y}(t) dt$  for all  $-\infty \leq a \leq b \leq \infty$ .

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*Solution.* Let  $t \in \mathbb{R}$ . Then

$$\begin{aligned}
 f_{X+Y}(t) &= \frac{d}{dt} \mathbf{P}(X + Y < t) = \frac{d}{dt} \iint_{\{x+y < t\}} f_{X,Y}(x,y) dx dy \\
 &= \frac{d}{dt} \iint_{\{x+y < t\}} f_X(x) f_Y(y) dx dy, && \text{since } X, Y \text{ are independent} \\
 &= \frac{d}{dt} \int_{y=-\infty}^{y=\infty} \int_{x=-\infty}^{x=t-y} f_X(x) f_Y(y) dx dy \\
 &= \int_{y=-\infty}^{y=\infty} f_X(t-y) f_Y(y) dy, && \text{by the Fundamental Theorem of Calculus} \\
 &= \frac{1}{\sqrt{2\pi}} \frac{1}{\sqrt{2\pi}} \int_{y=-\infty}^{y=\infty} e^{-(t-y)^2/2} e^{-y^2/2} dy, && \text{by definition of } X, Y \\
 &= \frac{1}{\sqrt{2\pi}} \frac{1}{\sqrt{2\pi}} \int_{y=-\infty}^{y=\infty} e^{-t^2/2 - y^2 + ty} dy = \frac{1}{\sqrt{2\pi}} e^{-t^2/4} \frac{1}{\sqrt{2\pi}} \int_{y=-\infty}^{y=\infty} e^{-(y-t/2)^2} dy \\
 &= \frac{1}{\sqrt{2\pi}} e^{-t^2/4} \frac{1}{\sqrt{2\pi}} \int_{y=-\infty}^{y=\infty} e^{-y^2} dy \\
 &= \frac{1}{2\sqrt{\pi}} e^{-t^2/4} \frac{1}{\sqrt{2\pi}} \int_{y=-\infty}^{y=\infty} e^{-y^2/2} dy, && \text{changing variables} \\
 &= \frac{1}{2\sqrt{\pi}} e^{-t^2/4}
 \end{aligned}$$

### 3. QUESTION 3

Suppose we have a Markov Chain  $(X_0, X_1, \dots)$  with state space  $\Omega = \{1, 2, 3, 4, 5\}$  and with the following transition matrix

$$P = \begin{pmatrix} 1/4 & 3/4 & 0 & 0 & 0 \\ 3/4 & 1/4 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 1/2 & 0 & 1/2 \\ 0 & 0 & 1/2 & 1/2 & 0 \end{pmatrix}.$$

Classify all states in the Markov chain as either transient or recurrent.

Is this Markov Chain irreducible? Prove your assertions.

*Solution.* State 1 is recurrent, since

$$\mathbf{P}_1(T_1 = \infty) = \mathbf{P}_1(2 = X_2 = X_3 = X_4 = \dots) = \lim_{n \rightarrow \infty} P(1, 2)(P(2, 2))^n = \lim_{n \rightarrow \infty} (3/4)(1/4)^n = 0.$$

State 2 is recurrent, since

$$\mathbf{P}_2(T_2 = \infty) = \mathbf{P}_2(1 = X_2 = X_3 = X_4 = \dots) = \lim_{n \rightarrow \infty} (P(1, 1))^n = \lim_{n \rightarrow \infty} (1/4)^n = 0.$$

State 3 is recurrent since  $\mathbf{P}_3(T_3 = 1) = 1$ , so  $\mathbf{P}_3(T_3 < \infty) = 1$ .

States 4 and 5 are transient, since

$$\mathbf{P}_4(T_4 = \infty) \geq \mathbf{P}_4(3 = X_2 = X_3 = X_4 = \dots) = P(4, 3) \lim_{n \rightarrow \infty} P(3, 3)^n = P(4, 3) > 0.$$

$$\mathbf{P}_5(T_5 = \infty) \geq \mathbf{P}_5(3 = X_2 = X_3 = X_4 = \dots) = P(5, 3) \lim_{n \rightarrow \infty} P(3, 3)^n = P(5, 3) > 0.$$

The Markov chain is not irreducible, since  $P$  is a block matrix of the form  $P = \begin{pmatrix} A & 0 \\ 0 & B \end{pmatrix}$  where  $A$  is  $2 \times 2$  and  $B$  is  $3 \times 3$ . So, for any  $n \geq 1$ ,  $P^n = \begin{pmatrix} A^n & 0 \\ 0 & B^n \end{pmatrix}$ . So,  $P^n(1, 3) = 0$  for any  $n \geq 1$ . So, the Markov chain is not irreducible.

#### 4. QUESTION 4

Let  $x, y$  be any states in a finite irreducible Markov chain. Show that  $\mathbb{E}_x T_y < \infty$ .

*Solution.* From Lemma 3.27 in the notes, there exists  $0 < \alpha < 1$  and  $j > 0$  such that, for any  $x, y \in \Omega$  and for any  $k > 0$ ,  $\mathbf{P}_x(T_y > kj) \leq \alpha^k$ . So,  $\mathbf{P}_x(T_y > kj) \leq \alpha^k$ . So, using Remark 2.23 in the notes,

$$\begin{aligned} \mathbb{E}_x T_y &= \sum_{k=1}^{\infty} \mathbf{P}_x(T_y \geq k) = \sum_{k=1}^{\infty} \sum_{j(k-1) < i \leq jk} \mathbf{P}_x(T_y \geq i) \\ &\leq \sum_{k=1}^{\infty} j \mathbf{P}_x(T_y > j(k-1)) \leq j \sum_{k=1}^{\infty} \alpha^{k-1} = j/(1-\alpha) < \infty. \end{aligned}$$

#### 5. QUESTION 5

Prove or disprove the following statement.

Every finite Markov chain on a nonempty state space has at least one recurrent state.

*Solution.* This statement is True. We argue by contradiction. Suppose every state is transient. That is, every  $y \in \Omega$  satisfies  $0 \leq \rho_{yy} < 1$ , where  $\rho_{yy} = \mathbf{P}_y(T_y < \infty)$ . As in the notes, let  $T_y^{(1)} := T_y$ , and for any  $k \geq 2$ , define a random variable  $T_y^{(k)} := \min\{n > T_y^{(k-1)} : X_n = y\}$ . That is,  $T_y^{(k)}$  is the  $k^{\text{th}}$  return time of the Markov chain. If  $k \geq 1$  is fixed, by the pigeonhole principle, at the  $|\Omega| \cdot k^{\text{th}}$  step of the Markov chain, the chain has returned to some state  $k$  times. That is, for any  $k \geq 1$  fixed, there exists  $y \in \Omega$  such that  $T_y^{(k)} < \infty$  with probability one. That is, for any fixed  $k \geq 1$ ,

$$1 = \mathbf{P}(\cup_{y \in \Omega} \{T_y^{(k)} < \infty\}).$$

Using the union bound,

$$1 \leq \sum_{y \in \Omega} \mathbf{P}(T_y^{(k)} < \infty).$$

From Proposition 3.21 in the notes,  $\mathbf{P}(T_y^{(k)} < \infty) = \rho_{yy}^k$ . So,

$$1 \leq \sum_{y \in \Omega} \rho_{yy}^k \leq |\Omega| (\max_{y \in \Omega} \rho_{yy})^k.$$

Since  $\Omega$  is finite,  $0 \leq \max_{y \in \Omega} \rho_{yy} < 1$ . So, if  $k > \frac{\log|\Omega|}{\log(1/\max_{y \in \Omega} \rho_{yy})}$ , we have  $1 < 1$ , a contradiction. The proof is complete.