

Please provide complete and well-written solutions to the following exercises.

Due February 21, in the discussion section.

## Homework 6

**Exercise 1.** Let  $\Omega = [0, 1]$ . Let  $\mathbf{P}$  be the uniform probability law on  $\Omega$ . Let  $X: [0, 1] \rightarrow \mathbf{R}$  be a random variable such that  $X(t) = t^2$  for all  $t \in [0, 1]$ . Let

$$\mathcal{A} = \{[0, 1/4), [1/4, 1/2), [1/2, 3/4), [3/4, 1]\}$$

Compute explicitly the function  $\mathbf{E}(X|\mathcal{A})$ . (It should be constant on each of the partition elements.) Draw the function  $\mathbf{E}(X|\mathcal{A})$  and compare it to a drawing of  $X$  itself.

Now, for every integer  $k > 1$ , let  $s = 2^{-k}$ , and let  $\mathcal{A}_k := \{[0, s), [s, 2s), [2s, 3s), \dots, [1-2s, 1-s), [1-s, 1]\}$ . Try to draw  $\mathbf{E}(X|\mathcal{A}_k)$ . Convince yourself of the following fact (you can prove it if you want, but you do not have to): for every  $t \in [0, 1]$

$$\lim_{k \rightarrow \infty} \mathbf{E}(X|\mathcal{A}_k)(t) = X(t).$$

The purpose of this exercise is to demonstrate that  $\mathbf{E}(X|\mathcal{A})$  is given by averaging  $X$  over each partition element, such that  $\mathbf{E}(X|\mathcal{A})$  is constant on each partition element of  $\mathcal{A}$ .

**Exercise 2.** Let  $X$  be a random variable with finite variance, and let  $t \in \mathbf{R}$ . Consider the function  $f: \mathbf{R} \rightarrow \mathbf{R}$  defined by  $f(t) = \mathbf{E}(X - t)^2$ . Show that the function  $f$  is uniquely minimized when  $t = \mathbf{E}X$ . That is,  $f(\mathbf{E}X) < f(t)$  for all  $t \in \mathbf{R}$  such that  $t \neq \mathbf{E}X$ . Put another way, setting  $t$  to be the mean of  $X$  minimizes the quantity  $\mathbf{E}(X - t)^2$  uniquely.

The conditional expectation, being a piecewise version of taking an average, has a similar property. Let  $A_1, \dots, A_k \subseteq \Omega$  such that  $A_i \cap A_j = \emptyset$  for all  $i, j \in \{1, \dots, k\}$  with  $i \neq j$ , and  $\cup_{i=1}^k A_i = \Omega$ . Write  $\mathcal{A} = \{A_1, \dots, A_k\}$ . By definition, for each  $1 \leq i \leq k$ ,  $\mathbf{E}(X|\mathcal{A})$  is constant on  $A_i$ . Now, let  $Y$  be any other random variable such that, for each  $1 \leq i \leq k$ ,  $Y$  is constant on  $A_i$ . Show that the quantity  $\mathbf{E}(X - Y)^2$  is uniquely minimized by such a  $Y$  only when  $Y = \mathbf{E}(X|\mathcal{A})$ .

**Exercise 3.** Let  $\Omega = [0, 1]$ . Let  $\mathbf{P}$  be the uniform probability law on  $\Omega$ . Let  $X: [0, 1] \rightarrow \mathbf{R}$  be a random variable such that  $X(t) = t^2$  for all  $t \in [0, 1]$ . For every integer  $k > 1$ , let  $s = 2^{-k}$ , let  $\mathcal{A}_k := \{[0, s), [s, 2s), [2s, 3s), \dots, [1-2s, 1-s), [1-s, 1]\}$ , and let  $M_k := \mathbf{E}(X|\mathcal{A}_k)$ . Show that the increments  $M_2 - M_1, M_3 - M_2, \dots$  are orthogonal in the following sense. For any  $i, j \geq 1$  with  $i \neq j$ ,

$$\mathbf{E}(M_{i+1} - M_i)(M_{j+1} - M_j) = 0.$$

This property is sometimes called **orthogonality of martingale increments**. This property holds for many martingales, but we will not prove this.

**Exercise 4.** Let  $X_0 = 0$ . Let  $(X_0, X_1, \dots)$  such that  $\mathbf{P}(X_i = 1) = \mathbf{P}(X_i = -1) = 1/2$  for all  $i \geq 1$ . For any  $n \geq 0$ , let  $Y_n = X_0 + \dots + X_n$ . So,  $(Y_0, Y_1, \dots)$  is a symmetric simple random walk on  $\mathbf{Z}$ . Show that  $Y_n^2 - n$  is a martingale (with respect to  $(X_0, X_1, \dots)$ ).

**Exercise 5.** Let  $1/2 < p < 1$ . Let  $(X_0, X_1, \dots)$  such that  $\mathbf{P}(X_i = 1) = p$  and  $\mathbf{P}(X_i = -1) = 1 - p$  for all  $i \geq 1$ . For any  $n \geq 0$ , let  $Y_n = X_0 + \dots + X_n$ . Let  $T_0 = \min\{n \geq 1: Y_n = 0\}$ . Prove that  $\mathbf{P}_1(T_0 = \infty) > 0$ . Then, deduce that  $\mathbf{P}_0(T_0 = \infty) > 0$ . That is, there is a positive probability that the biased random walk never returns to 0, even though it started at 0.

**Exercise 6.** Let  $X_1, \dots$  be independent identically distributed random variables with  $\mathbf{P}(X_i = 1) = \mathbf{P}(X_i = -1) = 1/2$  for every  $i \geq 1$ . For any  $n \geq 1$ , let  $M_n := X_1 + \dots + X_n$ . Let  $M_0 = 0$ . For any  $n \geq 1$ , define

$$W_n := M_0 + \sum_{m=1}^n H_m(M_m - M_{m-1}).$$

Show that if you have an infinite amount of money, then you *can* make money by using the double-your-bet strategy in the game of coinflips (where if you bet  $\$d$ , then you win  $\$d$  with probability  $1/2$ , and you lose  $\$d$  with probability  $1/2$ ). For example, show that if you start by betting  $\$1$ , and if you keep doubling your bet until you win (which should define some betting strategy  $H_1, H_2, \dots$  and a stopping time  $T$ ), then  $\mathbf{E}W_T = 1$ , for a suitable stopping time  $T$ .