

170B Midterm 2 Solutions¹

1. QUESTION 1

Label the following statements as TRUE or FALSE. If the statement is true, explain your reasoning. If the statement is false, provide a counterexample and explain your reasoning.

(a) Let X be a random variable. Let $i := \sqrt{-1}$. Then $|\mathbf{E}e^{itX}| \leq 1$ for any $t \in \mathbf{R}$.

TRUE. $|\mathbf{E}e^{itX}| \leq \mathbf{E}|e^{itX}| = 1$, using $|e^{itx}| = \sqrt{\cos^2(tx) + \sin^2(tx)} = 1$ for any $t, x \in \mathbf{R}$.

(b) Suppose I am flipping a coin over and over again. For any positive integer n , let A_n be the event that the n^{th} coin flip is heads. Suppose $\mathbf{P}(A_n) = n^{-2}$, for any positive integer n . Let B be the event that infinitely many of the coin flips are heads. Then $\mathbf{P}(B) = 0$.

TRUE. This follows from the Borel-Cantelli Lemma, since $\sum_{n=1}^{\infty} \mathbf{P}(A_n) = \sum_{n=1}^{\infty} n^{-2} < \infty$, so $\mathbf{P}(B) = 0$.

(c) Let X_1, X_2, \dots be independent random variables. Let $\mu := \mathbf{E}X_1$ and let $\sigma := \text{var}(X_1)$. Assume $0 < \sigma < \infty$ and $-\infty < \mu < \infty$. Then, for any $t \in \mathbf{R}$,

$$\lim_{n \rightarrow \infty} \mathbf{P}\left(\frac{X_1 + \dots + X_n - n\mu}{\sigma\sqrt{n}} \leq t\right) = \int_{-\infty}^t e^{-x^2/2} \frac{dx}{\sqrt{2\pi}}.$$

FALSE. We made no mention of being identically distributed. To get a counterexample, let X_1 so that $\mathbf{P}(X_1 = 1) = \mathbf{P}(X_1 = -1) = 1/2$ and let $X_n = 0$ for all $n \geq 2$. (Constant functions are automatically independent of other random variables.) Then $\mu = 0, \sigma = 1$ and $\frac{X_1 + \dots + X_n - n\mu}{\sigma\sqrt{n}} = \frac{X_1}{\sqrt{n}}$. So,

$$\lim_{n \rightarrow \infty} \mathbf{P}\left(\frac{X_1 + \dots + X_n - n\mu}{\sigma\sqrt{n}} \leq t\right) = \begin{cases} 0 & \text{if } t < 0 \\ 1 & \text{if } t \geq 0. \end{cases}$$

Moreover, the Central Limit Theorem should divide by the square root of the variance, not the variance, so this statement is false in two ways.

2. QUESTION 2

Let X_1, X_2, \dots be independent, identically distributed random variables such that $\mathbf{E}|X_1| < \infty$ and $\text{var}(X_1) < \infty$. For any $n \geq 1$, define

$$Y_n := \frac{1}{n} \sum_{i=1}^n X_i^2.$$

Show that Y_1, Y_2, \dots converges in probability. Express the limit in terms of $\mathbf{E}X_1$ and $\text{var}(X_1)$.

Solution. Note that X_1^2, X_2^2, \dots are independent, identically distributed random variables (since X_1, X_2, \dots are as well). (For example, $\mathbf{P}(X_i^2 \leq t) = \mathbf{P}(-\sqrt{t} \leq X_i \leq \sqrt{t}) = \mathbf{P}(-\sqrt{t} \leq X_1 \leq \sqrt{t}) = \mathbf{P}(X_1^2 \leq t)$ for any $t \geq 0, i \geq 1$, where the middle inequality follows since X_1, X_2, \dots are identically distributed.) So, the Weak Law of Large Numbers says that Y_1, Y_2, \dots converges in probability to its mean. In this case, we have for any $n \geq 1$,

$$\mathbf{E}Y_n = \frac{n}{n} \mathbf{E}X_1^2 = \mathbf{E}X_1^2 = \mathbf{E}X_1^2 - (\mathbf{E}X_1)^2 + (\mathbf{E}X_1)^2 = \text{var}(X_1) + (\mathbf{E}X_1)^2.$$

So, Y_1, Y_2, \dots converges in probability to the constant random variable $\text{var}(X_1) + (\mathbf{E}X_1)^2$.

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3. QUESTION 3

Let X be a random variable. Let X_1, X_2, \dots be a sequence of random variables such that

$$\lim_{n \rightarrow \infty} \mathbf{E} |X_n - X|^4 = 0.$$

Prove that X_1, X_2, \dots converges in probability to X .

Solution. Let $\varepsilon > 0$. From Markov's Inequality,

$$0 \leq \mathbf{P}(|X_n - X| > \varepsilon) = \mathbf{P}(|X_n - X|^4 > \varepsilon^4) \leq \varepsilon^{-4} \mathbf{E} |X_n - X|^4.$$

So, letting $n \rightarrow \infty$ and using our assumption, we get $0 \leq \lim_{n \rightarrow \infty} \mathbf{P}(|X_n - X| > \varepsilon) \leq 0$. That is, $\lim_{n \rightarrow \infty} \mathbf{P}(|X_n - X| > \varepsilon) = 0$. Since this holds for any $\varepsilon > 0$, we conclude that X_1, X_2, \dots converges in probability to X .

4. QUESTION 4

Let $f, g: \mathbf{R} \rightarrow \mathbf{R}$. Recall that $(f * g)(t) = \int_{-\infty}^{\infty} f(x)g(t-x)dx$. Show that, for any $t \in \mathbf{R}$,

$$(f * g)(t) = (g * f)(t).$$

Solution.

$$\begin{aligned} (f * g)(t) &= \int_{-\infty}^{\infty} f(x)g(t-x)dx, && \text{by the definition of convolution} \\ &= - \int_{\infty}^{-\infty} f(t-u)g(u)du, && \text{changing variables with } u = t-x, \text{ so } du = -dx \\ &= \int_{-\infty}^{\infty} g(u)f(t-u)du, && \text{by a property of integrals} \\ &= (g * f)(t), && \text{by the definition of convolution} \end{aligned}$$

5. QUESTION 5

Let X, Y be independent random variables. Suppose X is uniformly distributed in $[0, 1]$. And suppose Y has density given by

$$f_Y(t) = \begin{cases} 0 & , \text{ if } t < 0 \\ t & , \text{ if } 0 \leq t \leq 1 \\ 2-t & , \text{ if } 1 \leq t \leq 2 \\ 0 & , \text{ if } t > 2. \end{cases}$$

Find the density of $X + Y$.

Solution 1. From Proposition 2.60 in the notes, $f_{X+Y} = f_X * f_Y$. So, we compute $f_X * f_Y$, by breaking into several cases. Using that $f_Y(t-x) = 0$ except when $0 \leq t-x \leq 2$, i.e. when $0 \geq x-t \geq -2$, i.e. when $t-2 \leq x \leq t$, we have $f_X(x)f_Y(t-x) = 0$ except when

$0 \leq x \leq 1$ and $t - 2 \leq x \leq t$.

$$\begin{aligned}
f_X * f_Y(t) &= \int_{-\infty}^{\infty} f_X(x)f_Y(t-x)dx \quad , \text{ by the definition of convolution} \\
&= \int_{x=0}^{x=1} f_Y(t-x)dx \quad , \text{ by the definition of } f_X \\
&= \begin{cases} \int_{x=0}^{x=1} 0dx & , \text{ if } t < 0 \\ \int_{x=0}^{x=t} f_Y(t-x)dx & , \text{ if } 0 \leq t \leq 1 \\ \int_{x=0}^{x=1} f_Y(t-x)dx & , \text{ if } 1 \leq t \leq 2 \\ \int_{x=t-2}^{x=1} f_Y(t-x)dx & , \text{ if } 2 \leq t \leq 3 \\ 0 & , \text{ if } t > 3. \end{cases}
\end{aligned}$$

So, using the definition of f_Y , we get

$$\begin{aligned}
f_X * f_Y(t) &= \begin{cases} \int_{x=0}^{x=1} 0dx & , \text{ if } t < 0 \\ \int_{x=0}^{x=t} (t-x)dx & , \text{ if } 0 \leq t \leq 1 \\ \int_{x=0}^{x=t-1} f_Y(t-x)dx + \int_{x=t-1}^{x=1} f_Y(t-x)dx & , \text{ if } 1 \leq t \leq 2 \\ \int_{x=t-2}^{x=1} (2-(t-x))dx & , \text{ if } 2 \leq t \leq 3 \\ 0 & , \text{ if } t > 3. \end{cases} \\
&= \begin{cases} 0 & , \text{ if } t < 0 \\ [tx - x^2/2]_{x=0}^{x=t} & , \text{ if } 0 \leq t \leq 1 \\ \int_{x=0}^{x=t-1} (2-(t-x))dx + \int_{x=t-1}^{x=1} (t-x)dx & , \text{ if } 1 \leq t \leq 2 \\ [(2-t)x + x^2/2]_{x=t-2}^{x=1} & , \text{ if } 2 \leq t \leq 3 \\ 0 & , \text{ if } t > 3. \end{cases} \\
&= \begin{cases} 0 & , \text{ if } t < 0 \\ t^2/2 & , \text{ if } 0 \leq t \leq 1 \\ [(2-t)x + x^2/2]_{x=0}^{x=t-1} + [tx - x^2/2]_{x=t-1}^{x=1} & , \text{ if } 1 \leq t \leq 2 \\ (2-t + 1/2 + (2-t)^2 - (2-t)^2/2) & , \text{ if } 2 \leq t \leq 3 \\ 0 & , \text{ if } t > 3. \end{cases} \\
&= \begin{cases} 0 & , \text{ if } t < 0 \\ t^2/2 & , \text{ if } 0 \leq t \leq 1 \\ (2-t)(t-1) + (t-1)^2/2 + t - 1/2 - t(t-1) + (t-1)^2/2 & , \text{ if } 1 \leq t \leq 2 \\ t^2/2 - 3t + 9/2 & , \text{ if } 2 \leq t \leq 3 \\ 0 & , \text{ if } t > 3. \end{cases} \\
&= \begin{cases} 0 & , \text{ if } t < 0 \\ t^2/2 & , \text{ if } 0 \leq t \leq 1 \\ -t^2 + 3t - 3/2 & , \text{ if } 1 \leq t \leq 2 \\ t^2/2 - 3t + 9/2 & , \text{ if } 2 \leq t \leq 3 \\ 0 & , \text{ if } t > 3. \end{cases}
\end{aligned}$$

Solution 2. From Proposition 2.60 in the notes, $f_{X+Y} = f_X * f_Y$. So, we compute $f_X * f_Y$, by breaking into several cases. From the previous problem, $f_X * f_Y = f_Y * f_X$, so we compute $f_Y * f_X$. Note that $f_X(t-x) = 1$ when $0 \leq t-x \leq 1$, and $f_X(t-x) = 0$ otherwise. So, $f_X(t-x) = 1$ when $0 \geq x-t \geq -1$, i.e. when $-1+t \leq x \leq t$. These inequalities constrain the integral defining $f_Y * f_X(t)$ as follows.

$$\begin{aligned} f_Y * f_X(t) &= \int_{-\infty}^{\infty} f_Y(x)f_X(t-x)dx && , \text{ by the definition of convolution} \\ &= \int_{x=t-1}^{x=t} f_Y(x)dx && , \text{ by the definition of } f_X \end{aligned}$$

Substituting now the (piecewise) definition of f_Y ,

$$\begin{aligned} f_Y * f_X(t) &= \begin{cases} \int_{x=t-1}^{x=t} 0dx & , \text{ if } t < 0 \\ \int_{x=0}^{x=t} xdx & , \text{ if } 0 \leq t \leq 1 \\ \int_{x=t-1}^{x=1} xdx + \int_{x=1}^{x=t} (2-x)dx & , \text{ if } 1 \leq t \leq 2 \\ \int_{x=t-1}^{x=2} (2-x)dx & , \text{ if } 2 \leq t \leq 3 \\ 0 & , \text{ if } t > 3. \end{cases} \\ &= \begin{cases} 0 & , \text{ if } t < 0 \\ t^2/2 & , \text{ if } 0 \leq t \leq 1 \\ (1 - (t-1)^2)/2 + 2(t-1) - (t^2-1)/2 & , \text{ if } 1 \leq t \leq 2 \\ 2(2 - (t-1)) - (4 - (t-1)^2)/2 & , \text{ if } 2 \leq t \leq 3 \\ 0 & , \text{ if } t > 3. \end{cases} \\ &= \begin{cases} 0 & , \text{ if } t < 0 \\ t^2/2 & , \text{ if } 0 \leq t \leq 1 \\ -t^2 + 3t - 3/2 & , \text{ if } 1 \leq t \leq 2 \\ t^2/2 - 3t + 9/2 & , \text{ if } 2 \leq t \leq 3 \\ 0 & , \text{ if } t > 3. \end{cases} \end{aligned}$$